

POISSON-GAUSSIAN NOISE PARAMETER ESTIMATION IN FLUORESCENCE MICROSCOPY IMAGING

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INTRODUCTION

State of the art:

✗ Non-optimality (unknown parameters not jointly estimated)

Proposed method:

✓ Iterative, flexible algorithm producing reliable results

PROBLEM

Poisson-Gaussian model:

$$R_{s,t} = \alpha Q_{s,t} + W_{s,t}$$

$$\begin{aligned} Q_{s,t} &\sim \mathcal{P}(u_s e^{-k_s t}) \\ W_{s,t} &\sim \mathcal{N}(c, \sigma^2) \end{aligned}$$

$u = (u_s)_{1 \leq s \leq S} \in [0, +\infty]^S$ - “true” signal
 $k = (k_s)_{1 \leq s \leq S} \in [0, +\infty]^S$ - bleaching rate
 c, σ^2 - mean & variance of Gaussian noise
 $\alpha \in \mathbb{R}$ - Poisson noise scaling parameter

Problem formulation:

$$\text{Find } \hat{\theta} = (\hat{u}, \hat{k}, \hat{\alpha}, \hat{c}, \hat{\sigma}^2)$$

$\theta = (u, k, \alpha, c, \sigma^2)$ - unknown parameters
 $r = (r_{s,t})_{1 \leq s \leq S, 1 \leq t \leq T}$ - observed vector
 $R = (R_{s,t})_{1 \leq s \leq S, 1 \leq t \leq T}$ - random field
 t - time index and s - space index

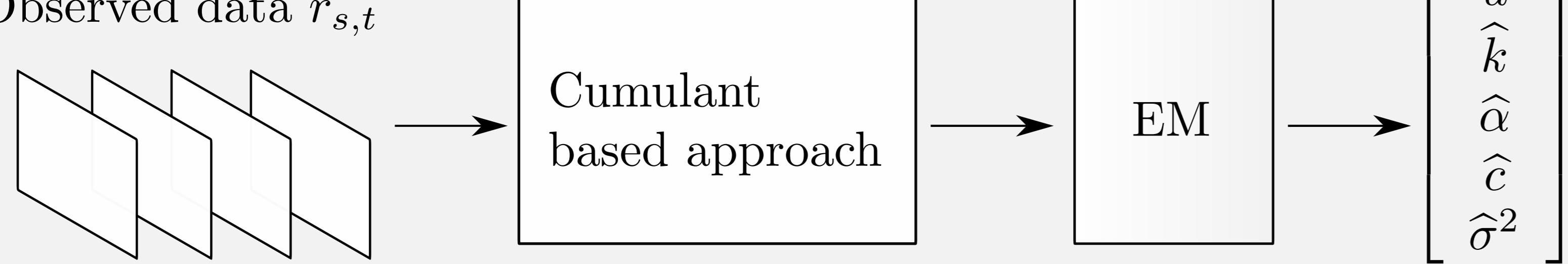
Challenge:

Find $2S + 3$ parameters with ST measurements

Assumptions:
 u - deterministic
 Q and W - mutually independent
Elements of W (resp Q) - independent

APPROACH

Observed data $r_{s,t}$



CUMULANT BASED INITIALIZATION

Cumulant of order n :

$$\kappa_n[R_{s,t}] = \alpha^n \kappa_n[Q_{s,t}] + \kappa_n[W_{s,t}] \quad \text{mean: } \kappa_1[R_{s,t}] = \alpha e^{-k_s t} u_s + c$$

$$\text{variance: } \kappa_2[R_{s,t}] = \alpha^2 e^{-2k_s t} u_s + \sigma^2$$

Observation 1: $R_{s,t} = a_s e^{-k_s t} + c + E_{s,t}$, $(a_s = \alpha u_s)_{1 \leq s \leq S}$

$(E_{s,t})_{1 \leq s \leq S, 1 \leq t \leq T}$ - independent zero-mean random variables

$$(\hat{a}, \hat{k}, \hat{c}) = \underset{a, k, c}{\operatorname{argmin}} \sum_{s=1}^S \sum_{t=1}^T (r_{s,t} - c - a_s e^{-k_s t})^2$$

Alternating minimization approach:

$$\text{Initialization: } c^{(0)} = \min\left\{\frac{1}{T} \sum_{t=1}^T r_{s,t}, 1 \leq s \leq S\right\}$$

For $n = 1 \dots N$

For $s = 1 \dots S$

$$(a_s^{(n)}, k_s^{(n)}) = \underset{a_s, k_s \geq 0}{\operatorname{argmin}} \sum_{t=1}^T (r_{s,t} - c^{(n-1)} - a_s e^{-k_s t})^2$$

$$c^{(n)} = \sum_{s=1}^S \sum_{t=1}^T (r_{s,t} - a_s^{(n)} e^{-k_s^{(n)} t}) / ST$$

$$\hat{a} = a^{(N)}, \hat{k} = k^{(N)}, \hat{c} = c^{(N)}$$

Observation 2: $E[(R_{s,t} - E[R_{s,t}])^2] = \alpha a_s e^{-k_s t} + \sigma^2$

~ Least squares estimator for α

~ $\sim (\forall s \in \{1, \dots, S\}) \quad \hat{u}_s = \frac{\hat{a}_s}{\hat{k}}$

~ $\sim \hat{\sigma}^2 = \left(\sum_{(s,t) \in \mathbb{I}} (v_{s,t} - \hat{a}_s \hat{k})^2 \right) / \text{card}(\mathbb{I}),$

where $v_{s,t} = (r_{s,t} - \hat{a}_s e^{-\hat{k}_s t} - \hat{c})^2$, \mathbb{I} - set for imposing positivity

EM

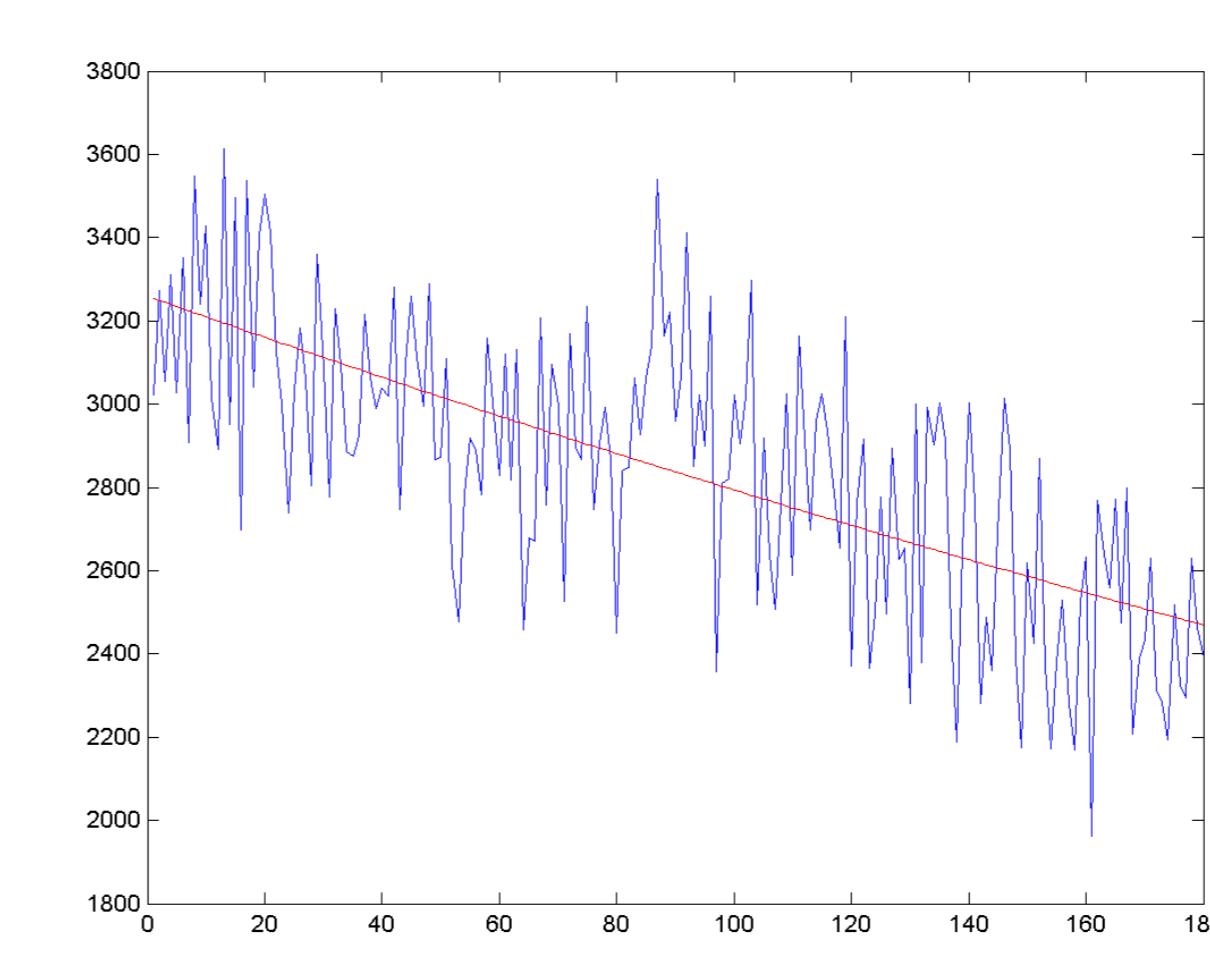
$$\text{Iterative algorithm: } (\forall n \in \mathbb{N}) \quad \theta^{(n+1)} = \underset{\theta}{\operatorname{argmax}} \quad J(\theta | \theta^{(n)})$$

$$J(\theta | \theta^{(n)}) = \mathbb{E}_{Q|R=r, \theta^{(n)}} [\ln p_{R,Q}(R, Q | \theta)]$$

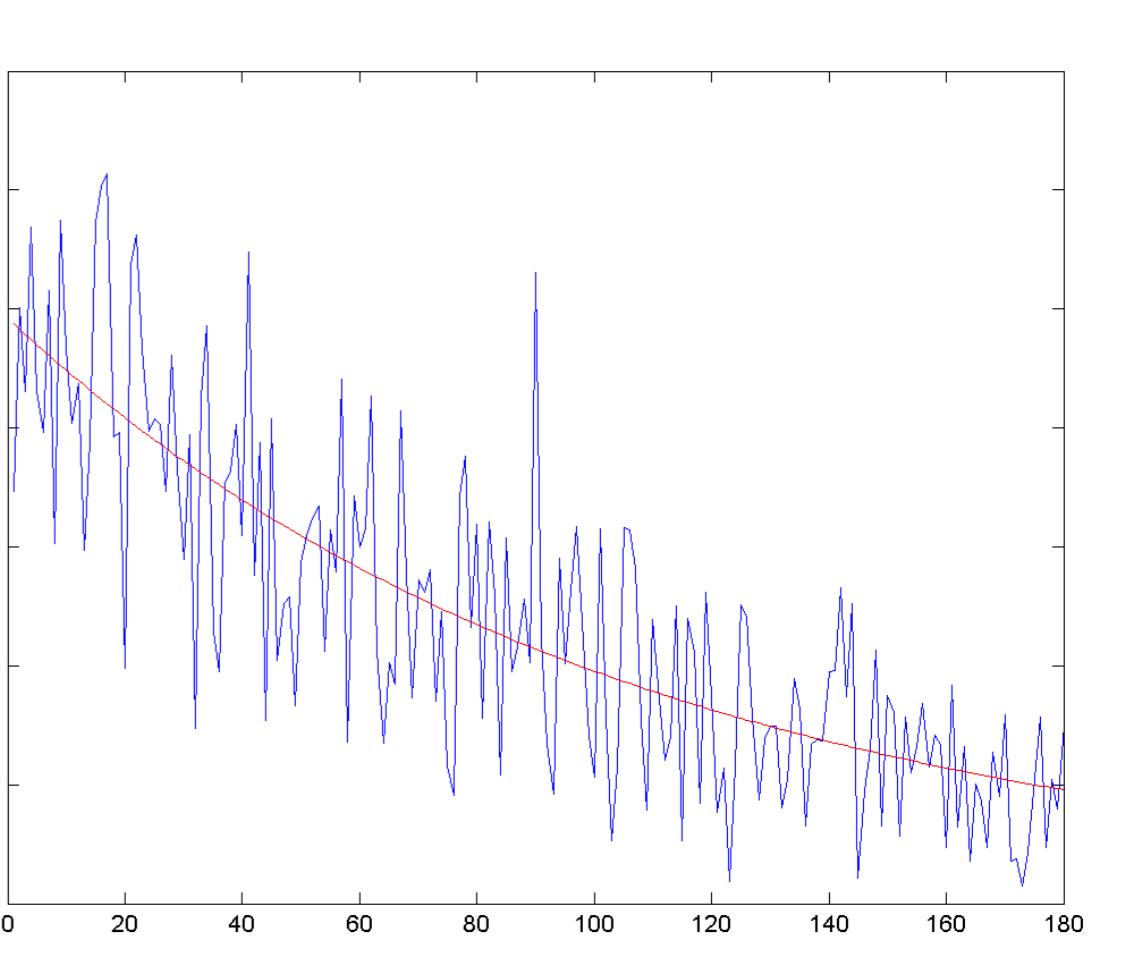
$p_{R,Q}(R, Q | \theta)$ - mixed continuous-discrete probability measure

- Main difficulties: computation of $\mathbb{E}_{Q|R=r, \theta^{(n)}}[Q_{s,t}]$ and $\mathbb{E}_{Q|R=r, \theta^{(n)}}[Q_{s,t}^2]$
- EM step based on [1]

TIME VARIATIONS

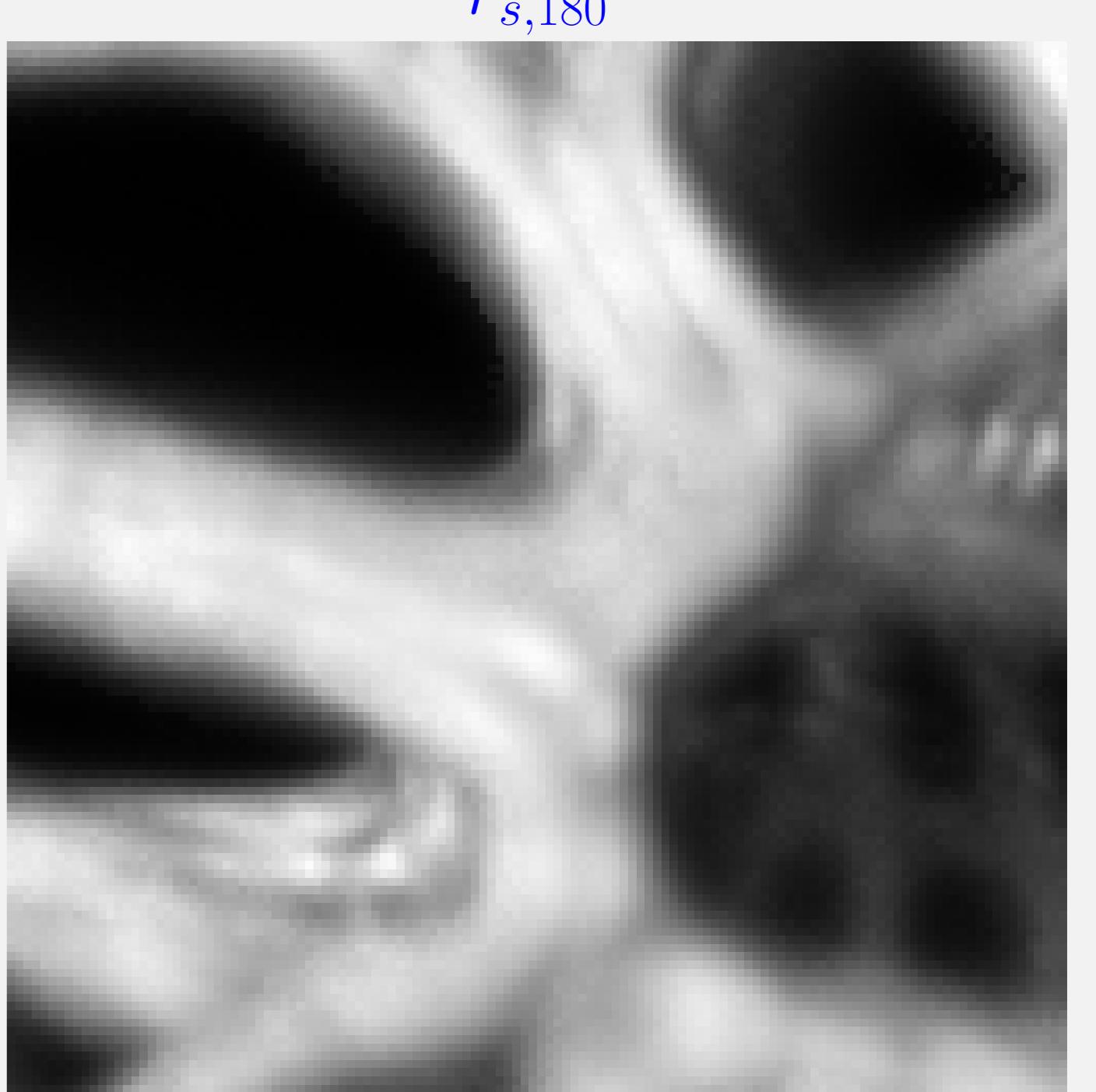
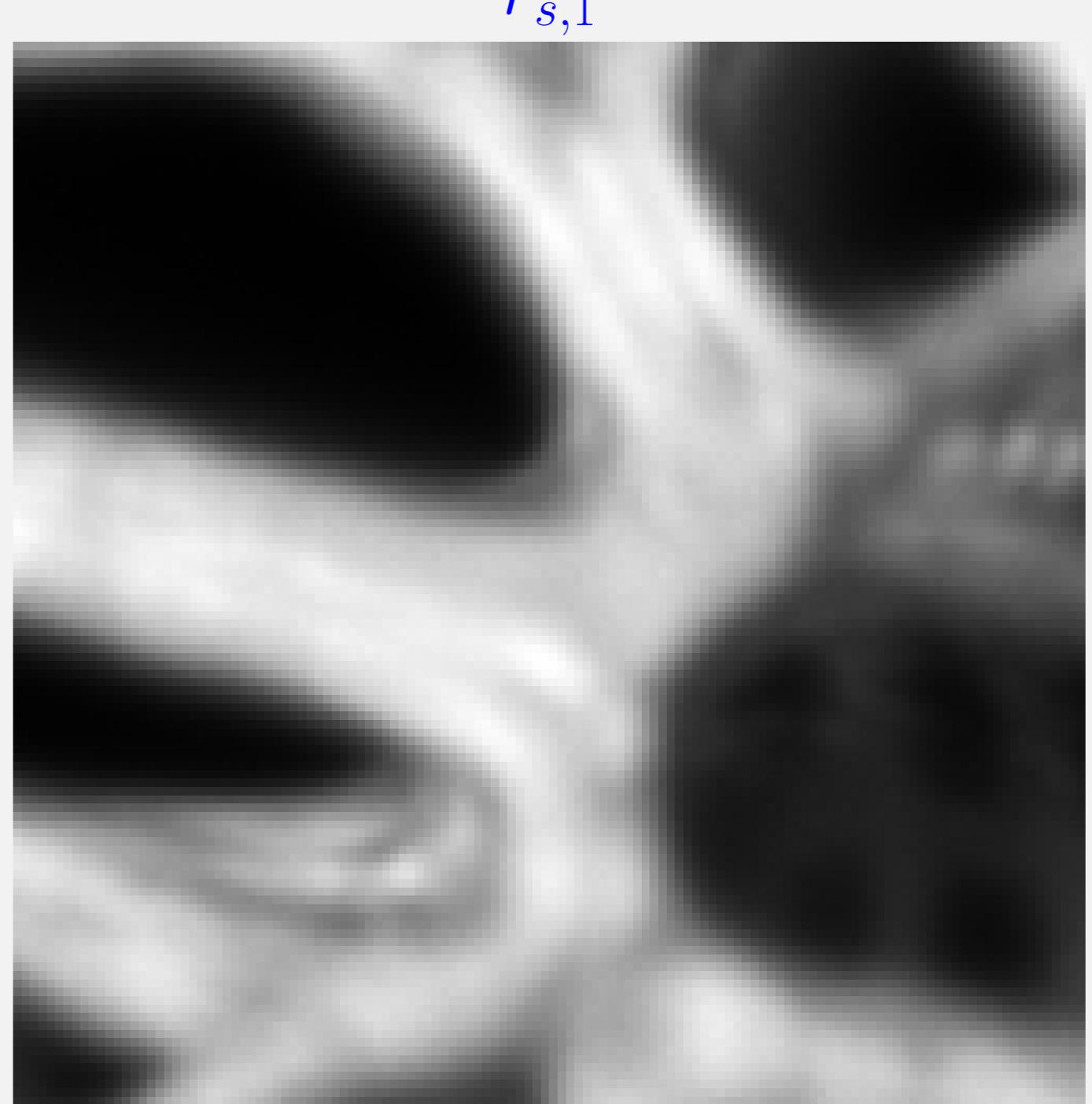
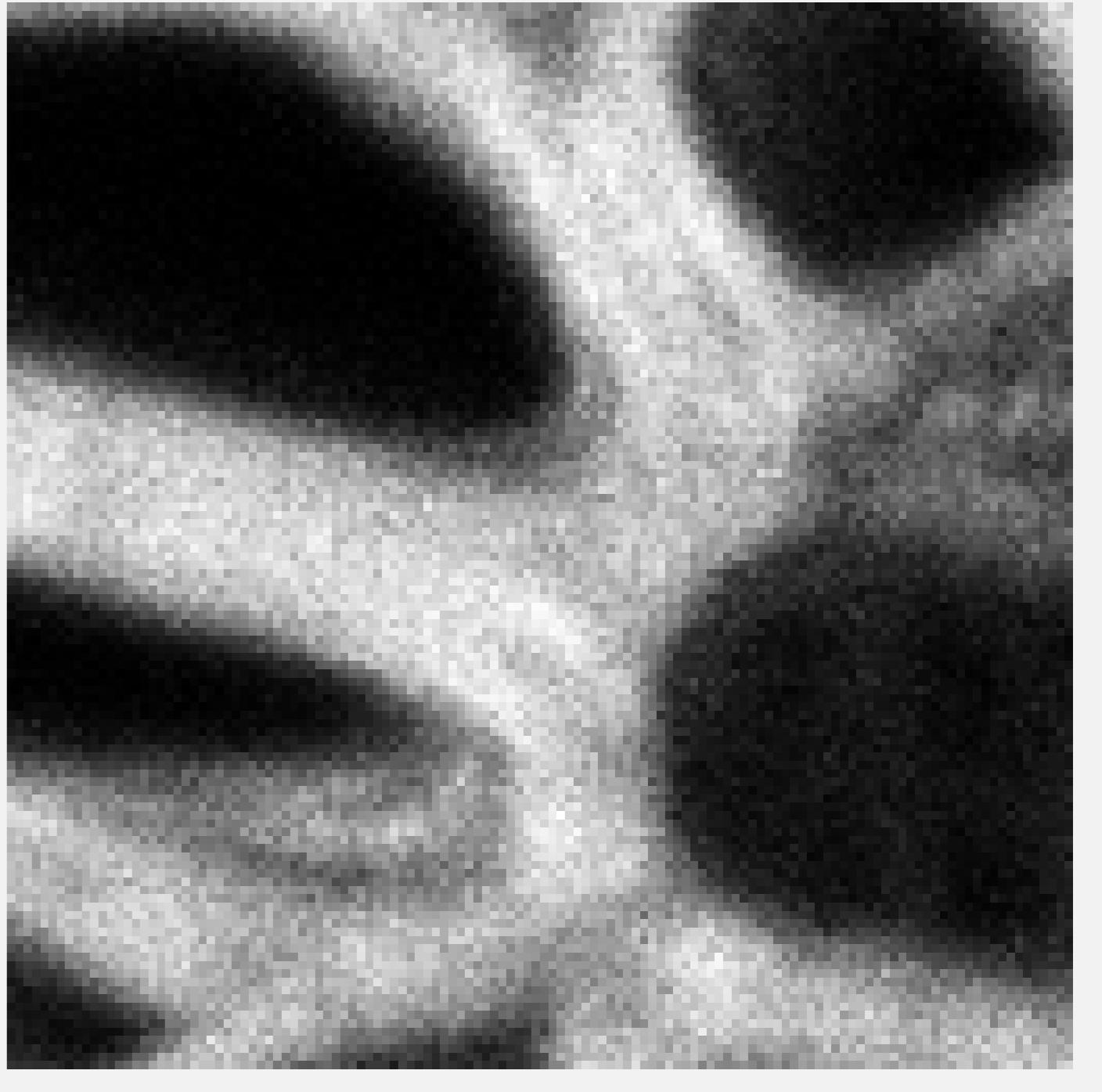
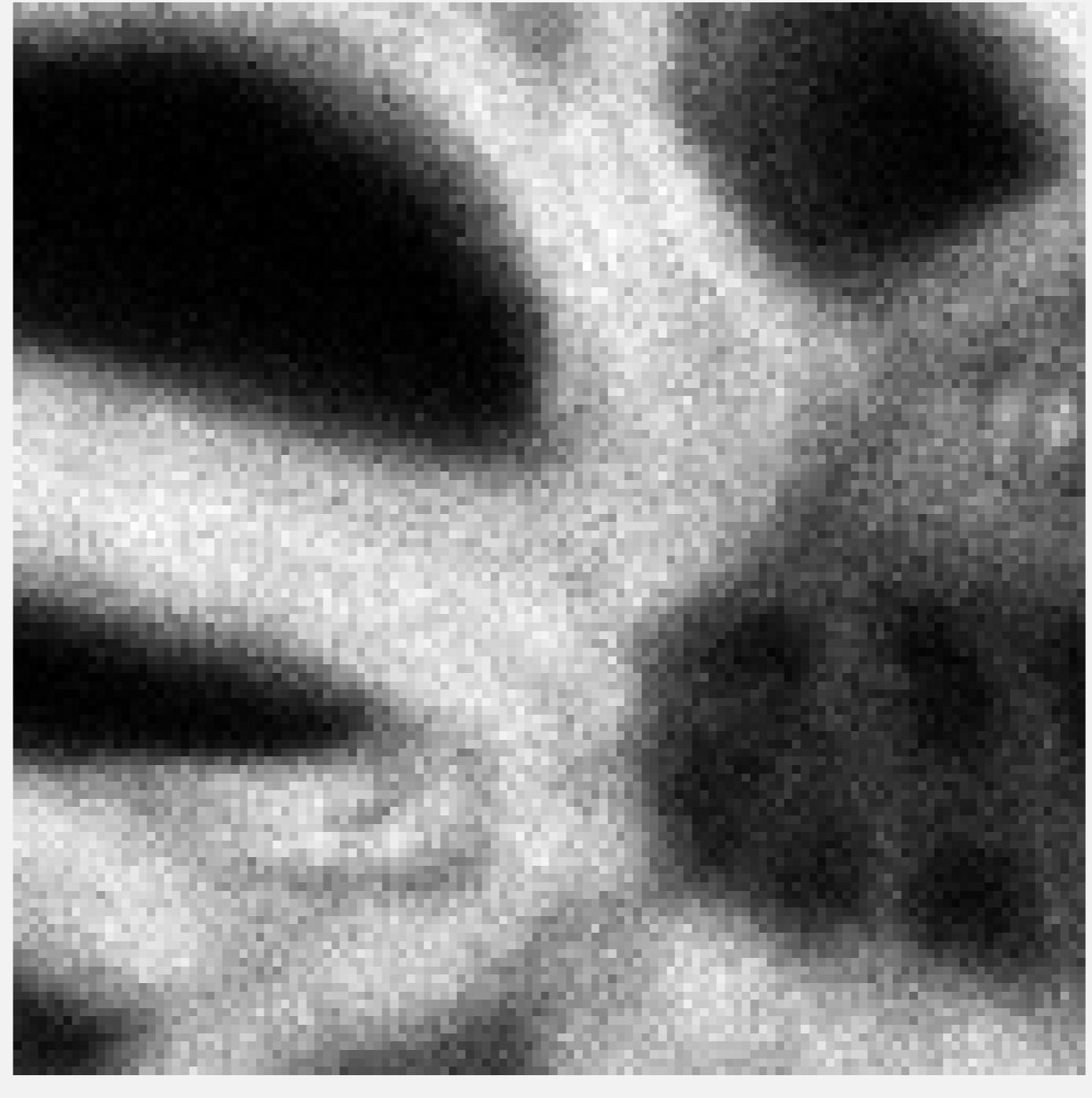


$$25.8 \times 126 e^{-5.7 \times 10^{-7} t} + 8$$



$$25.8 \times 19 e^{-3.6 \times 10^{-6} t} + 8$$

VISUAL RESULTS



$$\frac{1}{T} \sum_{t=1}^T r_{s,t}$$

$$\text{Reconstructed } \hat{u}_s$$

Images from macro confocal laser scanning microscope (Leica TCS-LSI)

Method	$\hat{\sigma}$		\hat{c}		$\hat{\alpha}$		SNR
	bias	std	bias	std	bias	std	
Init.	357.5	3.1	1.9	1.0	-0.3	0.4	39.5
EM	2.9	0.9	1.4	0.8	-0.3	0.4	39.7

Synthetic data results for $c = 10, \alpha = 30, \sigma^2 = 100$ and $T = 180$

References

- [1] A. Jezierska, C. Chaux, J.-C. Pesquet, and H. Talbot, An EM approach for Poisson-Gaussian noise modeling, European Signal Processing Conference (EUSIPCO), Barcelona, 29 August - 2 September 2011